Advances in data analysis using aggregated data

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CFE-CMStatistics 2024, 16 December 2024





Motivation

Big data → small (symbolic) data

General statistical questions:

- How to summarise a complex & very large dataset in a compact manner while retaining maximal relevant information in original dataset?
- How to do statistical analysis using symbolic data? What properties do the estimators have?

Useful for: Data storage, computational efficiency, data privatisation, data with non-standard form

In this talk

- 1. Present a general framework for data analysis through summaries
- 2. Asymptotic results (Prosha's work)
- 3. Design of histogram aggregation functions (Hakiim's work)

A possible approach to modelling aggregated data

Asymptotic results

Design of aggregation functions

One possible approach to modelling aggregated data

(Beranger, Lin & Sisson, 2023)

Define $S=\pi(X_{1:n}):[\mathcal{X}]^n o\mathcal{S}$ such that $x_{1:n}\mapsto\pi(x_{1:n})$ then,

$$L(S|\theta) \propto \int_{x} g(S|x,\phi) L(x|\theta) dx$$

where

- $L(x|\theta)$ standard, classical data likelihood
- $g(S|x,\phi)$ explains mapping to S given classical data x
- ullet L(S| heta) new "symbolic" likelihood for parameters of classical model

Gist

Fitting the standard classical model, when the data are viewed only through *symbols* S

Example: No generative model $L(x|\theta)$

- $g(S|x, \phi) = g(S|\phi) \Rightarrow L(S|\theta) = g(S|\phi)$
- Directly modelling symbol ⇒ (Le Rademacher & Billard, 2011)

Random bin histogram

Assume some fixed k_1, \ldots, k_B

Aggregation:

$$S = \pi(X_{1:n}) : \mathbb{R}^{d \times n} \to S = \{(a_1, \dots, a_B) \in \mathbb{R}^B : a_1 \le \dots \le a_B\} \times \mathbb{N}$$
$$x_{1:n} \mapsto (x_{(k_1)}, \dots, x_{(k_B)}, n)$$

Likelihood

$$\mathcal{L}_n(s|\theta) = n! \prod_{b=1}^B f(s_b|\theta) \prod_{b=1}^{B+1} \frac{(F(s_b|\theta) - F(s_{b-1}|\theta))^{k_b - k_{b-1} - 1}}{(k_b - k_{b-1} - 1)!}.$$

Key points:

- When B=2, $k_1=I$ and $k_2=u$ with $I, u=1, \ldots, n; I \neq u$ \Longrightarrow random intervals.
- Can recover classical likelihood if $B = n \Longrightarrow L(s|\theta) = f(x|\theta)$.

A possible approach to modelling aggregated data

Asymptotic results

Design of aggregation functions

Convergence of summaries

Setting:

Take random intervals, i.e., random bin histogram with B=2, $k_1=l$, $k_2=u$, and aggregation function $\pi(X_{1:n})=(X_{(l)},X_{(u)})$.

Things to consider:

Conditions on the sequences $1 \le l_n \le u_n \le n$ are needed to ensure asymptotically nondegenerate intervals: $l_n/n \to l_0$ and $u_n/n \to u_0$.

Approach:

Order statistics can be obtained from quantiles of the empirical distribution function (van der Vaart, 1998)

Convergence of summaries

Let $Q \in \mathcal{P}$ be a continuous distribution with empirical measure μ_n

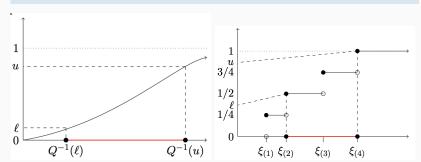
Interval-valued aggregation

Let
$$extbf{ extit{P}} = ig\{ (extit{I}, u) \in \mathbb{R}^2 : 0 < extit{I} \leq u < 1 ig\}$$
 and $\mathbb{R}^2_{\preceq} = ig\{ (a, b) \in \mathbb{R}^2 : a \leq b ig\}$

$$r: \mathcal{P} \times \boldsymbol{P} \to \mathbb{R}^2_{\leq}$$

 $(Q, (I, u)) \mapsto (Q^{-1}(I), Q^{-1}(u))$

Accordingly, the quantiles of μ_n are $r(\mu_n, (I, u)) = (X_{\lceil nI \rceil}, X_{\lceil nu \rceil})$.



Convergence of summaries

Convergence

The random interval $R_n(I, u) := r(\mu_n, (I, u))$ converges uniformly in probability to $R_{\infty}(I, u) := r(Q, (I, u))$.

Extensions

- Random rectangle (interval-valued aggregation to \mathbb{R}^d): R_n^d converges weakly to R_∞^d
- Random histograms: H_n^b converges uniformly almost surely to H_∞^b .
- \bullet Two distribution-valued aggregations with similar convergence properties

Convergence of the likelihood

Denote $S_n = \pi_n(X_1, \dots, X_n) = \pi(\mu_n)$, such that in the interval example $S_n(\omega) = R_n(I, u)(\omega)$

1 aggregate \Rightarrow the limit of the likelihood \mathcal{L}_n is determined by the limit of the sequence of densities f_{S_n} .

Suppose we fit the model P_{θ} , therefore

$$\star \ \mu_n o P_{ heta}$$
 weakly $\star \ S_n o \pi(P_{ heta})$ in probablity

Limit likelihood

For some true $\theta_0 \in \Theta$, we then get:

$$\mathcal{L}_{\infty}(\theta,\omega) = \lim_{n \to \infty} f_{S_n}(S_n(\omega))$$

$$= \lim_{n \to \infty} f_{\pi_n(X_1,...,X_n)}(S_n(\omega))$$

$$= \delta(\pi(P_{\theta_0}) - \pi(P_{\theta}))$$

Convergence of the likelihood

Convergence

- 1. The summary likelihood $\mathcal{L}_n \to \mathcal{L}_\infty$ uniformly in Θ in probability.
- 2. The MLE $\hat{\theta}_n \to \theta_0$ in probability and is a consistent estimator.

Extension

Convergence can be established for multiple data summaries (under some assumptions)

Summary

- ullet As we get more distributions, and data per distribution, the likelihood will consistently estimate θ_0 .
- Interest is now in the rate this happens (so we can design distributions with the most efficient rate).

A possible approach to modelling aggregated data

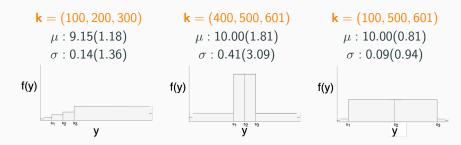
Asymptotic results

Design of aggregation functions

Illustrative example

Generate 1,001 samples from $\mathcal{N}(10,1)$.

Aggregation into 4 bin histograms with bins based on order statistics Fit the true model. Repeat 1,000 times.



Statistical decision theory

Let $\theta \in \Theta \subset \mathbb{R}^p$ be some unknown parameter of interest and $\mathbf{d} \in \mathcal{D}$ be some decision.

A loss function $L(\theta, \mathbf{d})$ measures the consequence of each decision \mathbf{d} , e.g., quadratic loss:

$$L(\theta, \mathbf{d}) = (\theta - \mathbf{d})^{\top} \mathbf{Q}(\theta - \mathbf{d})$$

This is not available since θ is unknown so we refer to the expected loss In the Bayesian framework, take some prior $p(\theta)$, the best belief about the distribution of θ is the posterior $p(\theta|s)$.

Posterior expected loss

$$ho(
ho(heta(heta|\mathbf{s}),\mathbf{d}) \equiv \mathbb{E}_{ heta|\mathbf{s}}\left[L(heta,\mathbf{d})
ight] = \int_{oldsymbol{\Theta}} L(heta,\mathbf{d})
ho(heta|\mathbf{s}) \mathrm{d} heta,$$

Statistical decision theory

Take $\mathbf{Q} = \mathbb{I}_p$, then $\mathbf{d}^* = \arg\min_{\mathbf{d} \in \mathcal{D}} \rho(p(\theta|\mathbf{y}), \mathbf{d}) = \mathbb{E}_{\theta|\mathbf{s}}[\theta]$ and

$$\rho(\pi(\theta|\mathbf{s}), \mathbf{d}^*) = \sum_{i=1}^{p} \mathbb{E}_{\theta_i|\mathbf{s}} \left[\left(\theta_i - \mathbb{E}_{\theta_i|\mathbf{s}} \left[\theta_i \right] \right)^2 \right] = \sum_{i=1}^{p} \mathbb{V}_{\theta_i|\mathbf{s}}(\theta_i).$$

Optimal design

An optimal symbolic data design minimises the minimised posterior expected loss (MPEL) function, $\mathbf{s}^* = \arg\min_{\mathbf{s}} \rho(p(\theta|\mathbf{s}), \mathbf{d}^*)$

Experiment: where to put the bins of a histogram?

Setup

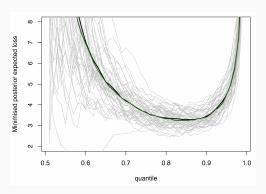
True model: $Y \sim \mathcal{N}(\mu = 50, \sigma = 17)$.

n = 201 observations;

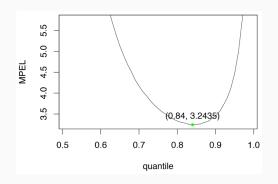
B = 2 (3 bins) with **symmetric** quantiles;

Compute the posterior Loss for varying quantiles;

Repeat 50 times (expensive!)



Experiment: where to put the bins of a histogram?



For a normal distribution, the suggests to use the 16 and 84% quantiles.

More (non-symmetric) quantiles:

$$B=2$$
 $B=3$ $B=4$ $B=5$ (0.14,0.85) (0.09,0.52,0.91) (0.07,0.32,0.74,0.95) (0.05,0.22,0.52,0.79,0.96)

Summary(!)

- Likelihood-based framework to fit model through summaries;
- Limit results ensuring the convergence of the summaries and the likelihood; Estimators have good properties: consistent!
- Bayesian framework for summary design.

THANK YOU

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